## PRIMES Circle Presentation

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## Definitions

Sample Space: The set of all possible outcomes, denoted by S.

**Event:** A subset of the sample space. Consists of possible outcomes of the experiment, denoted by A.

Union (AUB): outcomes that are either in A or in B or in both A and B

Intersection (AB): outcomes that are in both A and B, also denoted as A  $\cap$  B

**Complement: (A'):** outcomes in the sample space S that are not in A







#### **Axioms of Probability**

All probability of an event has to follow the 3 Axioms of Probability.

Axiom 1:  $0 \le P(E) \le 1$ 

**Axiom 2:** P(S) = 1

Axiom 3:

$$P\left(igcup_{i=1}^\infty E_i
ight) = \sum_{i=1}^\infty P(E_i)$$



#### **Conditional Probability**

Given two events E and F, The conditional probability is denoted by  $P(E \mid F)$ .

Essentially this conditional probability is the probability of both events happening with the new sample space F, so the definition and general formula for P (E | F) if P (F) > 0 is:

 $P(E|F) = \frac{P(EF)}{P(F)} \quad \Box \text{ both events happening} \\ \Box \text{ event F happening}$ 

P(F) has to be greater than 0 because a probability can't be negative, and P(E | F) would be undefined if P(F) equals 0.

#### **Multiplication Rule**

Gives the probability of the intersection of two events (both events happening).

P(EF) = P(F)P(E|F)

A generalization of the above equation gives the probability of the intersection of n events.

#### **Bayes' Formula**

 $P(F_j|E) = \frac{P(EF_j)}{P(E)}$ 

This formula is extremely useful because sometimes it is much easier to calculate the probability of an event when we know whether or not another event it depends on has occurred.

 $P(E_1 E_2 E_3 \cdots E_n) = P(E_1) P(E_2 | E_1) P(E_3 | E_1 E_2) \cdots P(E_n | E_1 \cdots E_{n-1})$ 

#### Independence

Two events E and F are independent if P(E | F) = P(E)In other words when P(EF) = P(E) \* P(F)

We can extend independence to more than two events. A set of events is independent if every subset of those events is independent.

 $P(E_{1'}, E_{2'}, \dots, E_{r'}) = P(E_{1'})P(E_{2'})\dots P(E_{r'})$ 

#### **Insurance Problem**

- Two types of people: those who are accident prone and those who are not
- Accident-prone person will have an accident in a 1-year period with probability 0.4, and a person who is not accident prone has probability 0.2
- 30% of the population is accident prone

Suppose that a new policyholder has an accident within a year of purchasing a policy. What is the probability that they are accident prone?

#### Problem

Bayes' Formula

 $P(A \mid B) = \frac{P(B \mid A) \cdot P(A)}{P(B)}$ 

Accident-prone person having an accident: 0.4 Non-accident-prone person having an accident: 0.2 Percent of population that is accident-prone: 0.3

A - the event that the policyholder is accident prone B - the event that the policyholder will have an accident within a year of purchasing the policy

$$P(A|B) = \frac{P(B|A) \cdot P(A)}{P(B)}$$
  
=  $\frac{(0.4)(0.3)}{P(B)}$   
=  $\frac{(0.4)(0.3)}{(0.4)(0.3) + (0.2)(0.7)}$   
=  $\frac{0.12}{0.26}$   
 $\approx 0.462$ 



## What is a Random Variable?

- A variable whose value is unknown, represented by a function that assigns values to each of an experiment's outcomes.
- Often designated by letters
- Can be classified as **discrete** or **continuous**
- Used in probability to quantify the outcome of a random experiment Ex: when flipping a coin, Heads=1 & Tails=0

#### Definitions

Discrete RV: A discrete random variable is a variable that can take any whole number values as outcomes of a random experiment.

 In real life, a discrete random variable could be represented as the number of passengers on a train or the number of defective computers out of a group of 100

It is defined by the probability mass function (PMF),  $p(a) = P{X = a}$ , which is always positive over a countable number of a.

There are various types of discrete random variables, and we are going to be discussing the Binomial, Geometric, and the Poisson RVs

#### **Expected Value & Variance**

**Expected Value:** the predicted value of a random variable

- calculated as the sum of all possible values each multiplied by the probability of its occurrence  $E[X] = \sum$
- essentially the weighted average of all the outcomes

Variance: degree of spread in the random variable

calculated as the average of squared deviations from the mean (denoted as  $\mu$ ), with formulas

$$Var(X) = E[(X - \mu)^2]$$
 OR  $Var(X) = E[X^2] - (E[X])^2$ 

xp(x)

x:p(x)>0

People often use expected value and variance to calculate the expected ROI for investments they make



**Definition:** the number of predicted "successes" in an experiment consisting of N trials – the experiment can only take on 2 values (success or failure)

**Parameters:** n = the number of trials, P = the probability of 1 trial succeeding, K = the number of desired successes

$$P(X = k) = \binom{n}{k} p^k (1-p)^{n-k}$$

Expected Value: n \* p

**Variance:** n \* p \* (1-p)



## **Binomial Example**

Let's say you buy some lottery tickets- 10 of the same kind. The probability that you get a winning ticket is 0.02. What is the probability that you get 3 winning tickets? If the tickets cost you \$5 each, and each winning ticket wins you \$100, was it worth it to buy those tickets- or were you scammed?

Let X be the number of winning tickets you get. We can apply the formula  $P(X = k) = \binom{n}{k} p^k (1-p)^{n-k}$ , letting n=10, k=3, and p=0.02, to find the desired probability:

$$P{X=3} = {}_{10}C_3 (0.03)^3 (1-0.03)^{10-3} = 0.0026.$$

Now, the expected value is simply n\*p, or 10\*0.03= 0.3 successes, and so your expected ROI is \$30 dollars-which means you're probably scammed :D

# Geometric RV

Definition: the number of predicted experiments that will be performed <u>before</u> reaching some number of "successes" (out of 2 options: success or failure)

**Parameter:** p = the probability of success for 1 trial, n (or sometimes k) = the number of the final (successful) trial.

Expected Value: 1/p - you need p \* 1/p to get 100% chance of success  $P(X = n) = p(1-p)^{n-1}$  $P(X > n) = (1-p)^n$ 

Variance:  $(1-p)/(p^2)$ 

#### **Geometric RV Example**

**Example Problem:** 3 out of 75 bulbs are defective, probability that the first defective light bulb is 6th?

#1: Define p, the probability that a light bulb is defective, and k, the number of trials needed to reach the "successful trial"

$$p = \frac{3}{75} = 0.04$$
$$k = 6$$

#2: Plug into formula

$$P(X = n) = p(1-p)^{n-1}$$

$$P(X = k) = p(1-p)^{k-1}$$

$$P(X = 6) = 0.04(1-0.04)^{6-1}$$

$$P(X = 6) = 0.04(0.96)^{5} = 0.0326$$



#### **Poisson RV**

Definition: the number of times an event is likely to occur over a specific interval (time, pages of a book, etc.)

 It is also a good approximation to binomial RVs for very large values of n and very small values of p

Parameter:  $\lambda$  = mean number of events in a given interval

Expected Value:  $\lambda$ 

Variance:  $\lambda$ 

 $P(x) = \frac{\lambda^2 e^{-x}}{|x|}$ 

#### **Poisson Example**

Let's say that the number of typos on any given slide of this presentation follows a Poison distribution of  $\lambda$ =0.5. What is the probability that there is at least one error on this slide

To solve this, we can let X be the Poisson RV, which is the number of errors on this slide.

Then, we can apply the formula and simplify to get

$$P\{X \ge 1\} = 1 - P\{X = 0\} = 1 - e^{-1/2} \approx .393$$

Note:  $e^{-\frac{1}{2}}$  is the simplified result of  $(0.5)^{0}(e^{-\frac{1}{2}})/0!$ .



## **Definitions:**

- A continuous random variable is a random variable that is defined over a range of values
  - $\circ~$  In real life, they are usually measurements (height, time required to run  $100^{+}$  meters, etc.)
- A continuous RV is defined by a **probability density function f(x) (PDF),** a non-negative function defined over all real numbers that functions similarly to the PMF in the discrete case
- Finding the probability of a continuous RV being within an interval (a,b) requires finding the **area under its probability curve** also known as taking an **integral** of the function between a and b.



### **Quick Example on Continuous RVs!**

Since we talked about expected value and variance earlier, and since the continuous RV is just another random variable, let's try to find E[X] and Var(X) of a continuous RV X, defined by

$$f(x) = \begin{cases} 2x & \text{if } 0 \le x \le 1\\ 0 & \text{otherwise} \end{cases}$$

Here are some formulas that we are going to use:

$$E[X] = \int_{-\infty}^{\infty} xf(x) dx \qquad \operatorname{Var}(X) = E[X^2] - (E[X])^2$$

#### **Quick Example on Continuous RVs!**

Now let's take a look at the solution:

$$E[X] = \int xf(x) \, dx$$
$$= \int_0^1 2x^2 \, dx$$
$$= \frac{2}{3}$$

#### **Quick Example on Continuous RVs!**

Now let's take a look at the solution:

**Solution.** We first compute  $E[X^2]$ .

$$E[X^{2}] = \int_{-\infty}^{\infty} x^{2} f(x) dx$$
$$= \int_{0}^{1} 2x^{3} dx$$
$$= \frac{1}{2}$$

Hence, since  $E[X] = \frac{2}{3}$ , we obtain

$$\operatorname{Var}(X) = \frac{1}{2} - \left(\frac{2}{3}\right)^2 = \frac{1}{18}$$



## Joint distribution functions

A joint distribution function is simply one that describes the probability of two or more random variables in one function.

For any two random variables X and Y, the joint probability distribution function (which describes the related probabilities of X & Y) is defined as

$$F(a,b) = P\{X \le a, Y \le b\} \quad -\infty < a, b < \infty$$

And their joint probability mass function is defined as

$$p(x, y) = P(X = x, Y = y)$$

## Joint RVs Example (Poisson)

Radioactive particles reach a Geiger counter (a machine that counts the number of radioactive particles) according to a Poisson process at a rate of  $\mu = 0.8$  particles per second.

What is the probability that the Geiger counter detects (exactly) 1 particle in the next second and 3 or more in the next 4 seconds?

Joint Probability: P(1 particle in (0, 1) AND 3 or more particles in (0, 4))

= P(1 particle in (0, 1) AND 2 or more particles in (1, 4)).

 $f_1 = \circ$  the number of particles in (0, 1) follows a  $Poisson(\mu = 0.8)$  distribution, and  $f_2 = \circ$  the number of particles in (1, 4) follows a  $Poisson(\mu = 2.4)$  distribution.

Plug in values!

$$f_1(1) \cdot (1 - f_2(0) - f_2(1)) = e^{-0.8} \frac{0.8^1}{1!} \cdot \left(1 - e^{-2.4} \frac{2.4^0}{0!} - e^{-2.4} \frac{2.4^1}{1!}\right) pprox .2486.$$



## Joint Distribution (Independent)

When computing a joint distribution, the two variables are independent of each other when

 $P(X \in A, Y \in B) = P(X \in A) \cdot P(Y \in B)$ 

In simple words, X and Y are independent if knowing the value of one does not change the value of the other.

If the variables are independent, the desired probabilities are multiplied, in accordance to the laws of conditional probability.



#### **Joint Distribution: Discrete**

- Suppose X and Y are two discrete random variables and that X takes values from  $\{x_1, x_2, \ldots, x_n\}$  and and Y takes values  $\{y_1, y_2, \ldots, y_m\}$
- The ordered pair (X, Y) take values in the product  $\{(x_1, y_1), (x_2, y_2), \dots, (x_n, y_m)\}$
- The joint probability mass function (joint pmf) of X and Y is the function  $p(x_i, y_j)$  giving the probability of the joint outcome  $X = x_i, Y = y_i$ .
- This is organized in a joint probability table
- The total probability must equal 1, represented by a double sum

$$\sum_{i=1}^n \sum_{j=1}^m p(x_i, y_j) = 1$$

$X \backslash Y$	$y_1$	$y_2$		$y_j$		$y_m$
$x_1$	$p(x_1,y_1)$	$p(x_1,y_2)$	•••	$p(x_1,y_j)$	•••	$p(x_1,y_m)$
$x_2$	$p(x_2,y_1)$	$p(x_2,y_2)$		$p(x_2, y_j)$		$p(x_2,y_m)$
		• • •		• • •		•••
			•••		•••	
$x_i$	$p(x_i,y_1)$	$p(x_i, y_2)$		$p(x_i,y_j)$	• • •	$p(x_i,y_m)$
		• • •	•••		• • •	
$x_n$	$p(x_n,y_1)$	$p(x_n, y_2)$		$p(x_n, y_j)$	•••	$p(x_n,y_m)$
					Ж	

## Joint Distribution: Discrete Example

• Suppose you roll two dice. Let X be the value on the first die and let T be the total on both dice. This would be represented in the joint probability table as follows

$X \backslash T$	2	3	4	5	6	7	8	9	10	11	12
1	1/36	1/36	1/36	1/36	1/36	1/36	0	0	0	0	0
2	0	1/36	1/36	1/36	1/36	1/36	1/36	0	0	0	0
3	0	0	1/36	1/36	1/36	1/36	1/36	1/36	0	0	0
4	0	0	0	1/36	1/36	1/36	1/36	1/36	1/36	0	0
5	0	0	0	0	1/36	1/36	1/36	1/36	1/36	1/36	0
6	0	0	0	0	0	1/36	1/36	1/36	1/36	1/36	1/36

## **Joint Distribution: Continuous**

- Continuous case: replace discrete sets of values by continuous intervals, and use the sums by integrals to compute
- If X takes values in [a, b] and Y takes values in [c, d] then the pair (X,Y) takes values in the product [a, b] × [c, d]
- The total probability must equal 1, expressed as a double integral:

$$\int_{c}^{d} \int_{a}^{b} f(x, y) \, dx \, dy = 1$$

 We will not be providing a worked example for this, as it involves a lot of calculus



#### **Markov Chain Sequence**

Consider an event with different **states** it could be in. A Markov chain is a model describing a **sequence** of events when the probability of each event depends **only** on the probability of the previous event.

A sequence of random variables  $X_0, X_1, ...,$  can be in states {0, 1, ..., M}.  $X_n$  is the state at time n, and the system is in state *i* at time *n* if  $X_n = i$ .

#### Example:

Any day in the town Lexington can be sunny, cloudy, rainy, or windy. The sequence of days are  $X_{\mathcal{O}}, X_{\mathcal{I}}, \dots$ ; the states are sunny, cloudy, rainy, or windy; the date would be time n.

#### **Transitional Probabilities**

The sequence of random variables  $X_{\mathcal{O}}, X_{\mathcal{I}}, \dots$ , is said to form a Markov chain if each time the system is in state i, there is a fixed probability,  $P_{j}$ , that the system will next be in state *j*.

The values  $P_{ij}$ ,  $0 \le i \le M$ ,  $0 \le j \le N$  are called *transition probabilities* of a markov chain if they satisfy:

1.  $P_{ij} \ge 0$  $\leftarrow$  The probability of being in state j if currently in state i can not be negative2.  $\sum_{j=0}^{M} P_{ij} = 1$  $\leftarrow$  An event has 100% probability of transitioning into the next event

3.  $i = 0, 1, ..., M \leftarrow An \text{ event could be in states } 0, 1, ..., M$ 

#### **Transitional Probabilities continued**

The transition probabilities can be represented as

$$P\{X_{n+1} = j \mid X_n = i, X_{n-1} = i_{n-1}, ..., X_1 = i_1, X_0 = i_0\} = P_{ij} \text{ for all } i_0, ..., i_{n-1}, i, j \in \mathbb{N}$$

#### Example:

In Lexington, if today is sunny, tomorrow has 0.5 chance of being sunny, 0.15 chance of being cloudy, 0.1 chance of being rainy, and 0.25 chance of being windy.

- If today is cloudy, tomorrow has 0.2 chance of being sunny, 0.15 chance of being cloudy, 0.5 chance of being rainy, and 0.15 chance of being windy.
- If today is rainy, tomorrow has 0.25 chance of being sunny, 0.5 chance of being cloudy, 0 chance of being rainy, and 0.25 chance of being windy.
- If today is windy, tomorrow has 0.6 chance of being sunny, 0.2 chance of being cloudy, 0 chance of being rainy, and 0.2 chance of being windy.

#### **Transition Matrix**

However, it is more convenient to arrange the transition probabilities  $P_{ij}$  in a matrix, and this matrix is called the transition probability matrix.

$$\begin{bmatrix} P_{00} & P_{01} & \cdots & P_{0M} \\ P_{10} & P_{11} & \cdots & P_{1M} \\ \vdots & & & \\ \vdots & & & \\ P_{M0} & P_{M1} & \cdots & P_{MM} \end{bmatrix}$$

For example, the transition matrix of weather in Lexington is:

day		Weather of next day							
ent		Sunny	Cloudy	Rainy	Windy				
curi	Sunny	0.5	0.15	0.1	0.25				
er of	Cloudy	0.2	0.15	0.5	0.15				
Weathe	Rainy	0.25	0.5	0	0.25	1			
	Windy	0.6	0.2	0	0.2				
		-				-			

\*Notice how each row adds up to 1

#### **Transformation Matrix**

The transition matrix can be used as a transformation matrix. The transformation matrix can be used in many different ways.

For example, the weather in Lexington can be represented as

$$T = \begin{bmatrix} 0.5 & 0.15 & 0.1 & 0.25 \\ 0.2 & 0.15 & 0.5 & 0.15 \\ 0.25 & 0.5 & 0 & 0.25 \\ 0.6 & 0.2 & 0 & 0.2 \end{bmatrix}$$

The probability of weather in Lexington in *n* days is:

$$\left[\begin{array}{rrrr}1&0&0&0\\0&1&0&0\\0&0&1&0\\0&0&0&1\end{array}\right]T'$$

#### **Transformation Matrix continued**

Continuing with our example, the probability of each weather in 2 days (if day 1 is sunny) is:

			ſ	0.5	0.15	0.1	0.25	וו	0.5	0.15	0.1	0.25	1
[ 1	0	0	010	0.2	0.15	0.5	0.15		0.2	0.15	0.5	0.15	1
		0	010	0.25	0.5	0	0.25	1.1	0.25	0.5	0	0.25	1
				0.6	0.2	0	0.2		0.6	0.2	0	0.2	

 $= \begin{bmatrix} 0.455 & 0.1975 & 0.125 & 0.2225 \end{bmatrix}$ 

So there's 45.5% chance it will be sunny again, 19.75% it will be cloudy, 12.5% it will be rainy, and 22.25% it will be windy in 2 days.

#### **Calculations and Observations**

ſ	1	0	0	0	1	0.432	0.205	0.144	0.219
	0	1	0	0	73	0.399	0.195	0.201	0.205
	0	0	1	0	1=	0.401	0.256	0.119	0.224
l	0	0	0	1		0.434	0.217	0.126	0.223
٢	1	0	0	0	1	0.422	0.212	0.148	0.218 ]
	0	1	0	0	775	0.419	0.209	0.156	0.216
	0	0	1	0	1 =	0.419	0.219	0.143	0.219
ł	0	0	0	1		0.422	0.214	0.145	0.219
Г	1	0	0	0.	1	0.421	0.213	0.149	0.218 ]
	0	1	0	0	cr15	0.421	0.213	0.149	0.218
	0	0	1	0	1=	0.421	0.213	0.149	0.218
L	0	0	0	1		0.421	0.213	0.149	0.218
Г	1	0	0	0 .	1	0.421	0.213	0.149	0.218
	0	1	0	0	cm30	0.421	0.213	0.149	0.218
	0	0	1	0	$I^{ab} =$	0.421	0.213	0.149	0.218
L	0	0	0	1		0.421	0.213	0.149	0.218
Γ	1	0	0	0 ]		0.421	0.213	0.149	0.218
	0	1	0	0	T100	0.421	0.213	0.149	0.218
	0	0	1	0	1 =	0.421	0.213	0.149	0.218
	0	0	0	1		0.421	0.213	0.149	0.218

calculated using calculator (with rounding)

 notice how each row eventually becomes similar

 notice how the values doesn't really change after a large number of transformations

#### **Steady State**

As seen in the previous slide, the probabilities doesn't really change after a large number of transformations. Eventually the system reaches a state where further transformations wouldn't change the probabilities. This is called a **steady state**, and it gives the long-term probability of a system.

The steady state can also be calculated with precise math.

#### **Steady State continued**

Let S be the steady state and T be the transition matrix.

ST = S  $\Box$  further transformations does not change the steady state

Sumy Cloudy Rainy Windy  

$$\begin{bmatrix} a & b & c & (1-a-b-c) \end{bmatrix} \begin{bmatrix} 0.5 & 0.15 & 0.1 & 0.25 \\ 0.2 & 0.15 & 0.5 & 0.15 \\ 0.25 & 0.5 & 0 & 0.25 \\ 0.6 & 0.2 & 0 & 0.2 \end{bmatrix} = \begin{bmatrix} a & b & c & (1-a-b-c) \end{bmatrix}$$

$$\begin{cases} 0.5a + 0.2b + 0.25c + 0.6(1-a-b-c) = a \\ 0.15a + 0.15b + 0.5c + 0.2(1-a-b-c) = b \\ 0.1a + 0.5b + 0c + 0(1-a-b-c) = c \\ 0.25a + 0.15b + 0.25c + 0.2(1-a-b-c) = 1-a-b-c \end{bmatrix}$$
System of equations
$$\begin{cases} 1.1a + 0.4b + 0.35c = 0.6 \\ 0.05a + 1.05b - 0.3c = 0.2 \\ 0.1a + 0.5b - c = 0 \\ 1.05a + 0.95b + 1.05c = 0.8 \end{bmatrix}$$

#### **Steady State continued**

#### Put into matrix form:

 $\begin{bmatrix} 1.1 & 0.4 & 0.35 & 0.6 \\ 0.05 & 1.05 & -0.3 & 0.2 \\ 0.1 & 0.5 & -1 & 0 \\ 1.05 & 0.95 & 1.05 & 0.8 \end{bmatrix}$ 

#### Using RREF, a method used for matrix calculations:

 $\begin{bmatrix} 1.1 & 0.4 & 0.35 & 0.6\\ 0.05 & 1.05 & -0.3 & 0.2\\ 0.1 & 0.5 & -1 & 0\\ 1.05 & 0.95 & 1.05 & 0.8 \end{bmatrix} = \begin{bmatrix} 1 & 0 & 0 & 0.4208\\ 0 & 1 & 0 & 0.2129\\ 0 & 0 & 1 & 0.1485\\ 0 & 0 & 0 & 0 \end{bmatrix}$ 

 $S = \begin{bmatrix} 0.4208 & 0.2129 & 0.1485 & 0.2178 \end{bmatrix}$ 

On any given day, Lexington has 42.08% chance of being sunny, 21.29% of being cloudy, 14.85% of being rainy, and 21.78% of being windy.

